Impact of age and sex on sudden cardiovascular death following myocardial infarction

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Objective: To evaluate and compare the risk of sudden cardiovascular death (SCD) and non-SCD after myocardial infarction (MI) associated with age and sex.

Design: Cohort study of patients admitted with an enzyme verified acute MI and discharged alive. Patients were followed up for up to four years.

Patients: 5983 consecutive hospital survivors of acute MI were enrolled in the TRACE (trandolapril cardiac evaluation) registry from 1990–92. Four age groups were prespecified: < 56, 56–65, 66–75, and > 76 years.

Main outcome measures: SCD was defined as cardiovascular death within one hour of onset of symptoms.

Results: There were 536 SCD and 725 non-SCD. SCD mortality was 4.8% in the youngest and 15.7% in the oldest age groups. Non-SCD mortality was 3.5% and 25%, respectively. The ratio of SCD to non-SCD mortality varied from 1.44 in the youngest (< 56 years) to 0.55 in the oldest patients (> 76 years).

Age significantly increased both SCD and non-SCD risk (p < 0.0001), but the increase in non-SCD risk was 40% higher (p < 0.0001). Male sex was associated with increased risk of SCD independently of age (risk ratio 1.34, p < 0.005). However, the absolute three year probability of SCD among women older than 66 years exceeded 10%.

Conclusions: Compared with non-SCD the risk of SCD is relatively highest in the younger age groups, but the absolute risk of SCD is much higher among the upper age groups than the younger. The risk of SCD was slightly lower in women but not enough to warrant a different treatment strategy.
as a percentage can be approximated by multiplying the wall motion index by 30.

**Mortality**

Survival status among all patients screened was obtained two years after screening of the last patient on 15 July 1994 (median follow up 32 months). By making use of the Civil Registration System of Denmark we had to censor only 31 (28 non-Danish) patients at the time of hospital discharge, when they were last known to be alive.

**Sudden cardiovascular death**

SCD was defined as cardiovascular death within one hour of onset (or significant worsening) of symptoms leading to death. A central mortality committee classified all deaths with respect to cause and mode using information from death certificates, police investigations, hospital records, and necropsy reports when available. One committee classified death in the 1749 patients who had been randomly assigned to treatment in the TRACE study, and a separate committee used the identical criteria to classify those not randomly assigned to treatment. In 32 random cases both committees determined the mode of death. Differences in classification were found in five cases with a \( \kappa \) statistic of 0.61, which indicates good agreement.

Firstly, the committee classified whether death was caused by cardiovascular disease. Whenever the evidence was inadequate or uncertain the cause of death was classified as unknown. When information was available, death was assumed to be cardiovascular unless proved otherwise. Secondly, cardiovascular death was classified as sudden or non-sudden death with particular emphasis on the time elapsed from onset of new symptoms to death. Only cardiovascular death with a period documented to be less than one hour was classified as SCD. If there was any doubt concerning the elapsed time, the death was classified as non-SCD. Deaths of patients who died during their sleep, without any preceding symptoms and known to have been alive within 12 hours before their death, were classified as SCD. Deaths of patients admitted with cardiac arrest but who died after a period in coma following resuscitation were classified as SCD.

**Statistical methods**

The patient population was divided into four prespecified age groups: < 56, 56–65, 66–75, and \( \geq 76 \) years. There were more than 1000 participants in each interval.

In univariate and multivariate studies of a single mode of death, Cox proportional hazard models were used. Variables with \( p < 0.05 \) in the models were considered to contribute independently. Since all patients in the analyses were required to have been discharged alive, and since lifetime was defined as time from MI to death or censoring, methods accounting for delayed entry (day of discharge) were applied.

The assumptions of the Cox model (possible interaction, and proportional hazards and log linearity of effects of quantitative covariates) were tested both graphically and numerically where appropriate.

The ratio between the risk of SCD and that of non-SCD associated with a given risk factor was estimated in a Cox model.

When evaluating other than all cause mortality, the particular mode of death was defined as the event while patients experiencing any other mode of death were censored. As the Kaplan-Meier estimator is not applicable to estimating probabilities of particular modes of death in a competing risk model, such probabilities were estimated as cumulative incidence functions. Cause specific hazards were statistically compared by the log rank test.

All analyses were performed using the SAS software (Statistical Analysis System, Cary, North Carolina, USA) version 6.12. The cumulative incidence functions were estimated from the output of the SAS PHREG procedure using a custom built program. The risk ratio denotes the hazard ratio.

**RESULTS**

Of the 6676 patients enrolled in the TRACE registry, 5983 were discharged alive from hospital and evaluated in this study. During follow up until June 1994, 1659 patients died. Among the deaths, 1261 were classified to be due to cardiovascular causes, 256 were classified to be due to non-cardiovascular causes, and 142 were unclassifiable because of inadequate information. Of the cardiovascular deaths 536 were classified as SCD and 725 as non-SCD.

The cumulative incidence of SCD in the entire population was 5.3%, 7.4%, and 9.4% at one, two, and three years, respectively. The risk of both SCD and non-SCD increased significantly with increasing age (fig 1). The risk ratio associated with a 10 years increase in age (independently of sex) was 1.56 and 2.13 for SCD and non-SCD, respectively (\( p < 0.0001 \)) (table 3). The increase in non-SCD associated with increasing age was 40% higher than the increase in SCD risk (\( p < 0.0001 \)).

Table 4 gives the three year mortality from SCD and non-SCD in each age group. The ratio between risk of SCD and

### Table 1 Baseline characteristics of 5983 myocardial infarction survivors according to age group

<table>
<thead>
<tr>
<th>Age group (years)</th>
<th>All (n=5983)</th>
<th>&lt;56 (n=1120)</th>
<th>56–65 (n=1531)</th>
<th>66–75 (n=2001)</th>
<th>&gt;76 (n=1331)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Men* [%]</td>
<td>69</td>
<td>84</td>
<td>75</td>
<td>68</td>
<td>52</td>
</tr>
<tr>
<td>Wall motion index**</td>
<td>[0.8–2.0]</td>
<td>[0.9–2.0]</td>
<td>[0.8–2.0]</td>
<td>[0.7–2.0]</td>
<td>[0.7–2.0]</td>
</tr>
<tr>
<td>Congestive failure* [%]</td>
<td>49</td>
<td>24</td>
<td>42</td>
<td>56</td>
<td>69</td>
</tr>
<tr>
<td>Smokers* [%]</td>
<td>53</td>
<td>76</td>
<td>64</td>
<td>47</td>
<td>29</td>
</tr>
<tr>
<td>Hypertension* [%]</td>
<td>22</td>
<td>16</td>
<td>22</td>
<td>25</td>
<td>23</td>
</tr>
<tr>
<td>Diabetes* [%]</td>
<td>10</td>
<td>7</td>
<td>9</td>
<td>12</td>
<td>13</td>
</tr>
<tr>
<td>Previous myocardial infarction* [%]</td>
<td>23</td>
<td>15</td>
<td>22</td>
<td>27</td>
<td>25</td>
</tr>
<tr>
<td>Body mass index**</td>
<td>[20–33]</td>
<td>[20–34]</td>
<td>[20–33]</td>
<td>[20–32]</td>
<td>[19–32]</td>
</tr>
<tr>
<td>Thrombolytic* [%]</td>
<td>43</td>
<td>54</td>
<td>52</td>
<td>42</td>
<td>23</td>
</tr>
<tr>
<td>Ventricular fibrillation* [%]</td>
<td>5</td>
<td>6</td>
<td>5</td>
<td>5</td>
<td>3</td>
</tr>
<tr>
<td>Atrial fibrillation* [%]</td>
<td>19</td>
<td>6</td>
<td>13</td>
<td>23</td>
<td>32</td>
</tr>
</tbody>
</table>

\*p<0.001, \**p<0.0001 for no difference between the four age groups. Continuous variables are medians with 5th and 95th centiles. Discrete variables were compared by \( \chi^2 \) tests and continuous variables by rank sum tests.
risk of non-SCD decreased significantly with increasing age (table 4), from 1.44 in the youngest to 0.55 in the oldest (p < 0.001).

In a multivariate analysis, male sex was associated with an increased risk of SCD and all cause mortality (table 3), but male sex had no predictive value concerning non-SCD independently of age (table 3). The increase in SCD risk associated with male sex was 28% higher than the increase in non-SCD risk (p = 0.05).

Stratification of the analyses by sex showed that the age effect was similar in men and women, since no interaction was present. Figure 2 illustrates this, which shows the cause specific cumulative three year mortalities. Furthermore, fig 2 shows that the SCD risk of women lagged less than 10 years behind that of men.

Since the registry included patients both with and without randomly assigned treatment, the analyses were repeated with inclusion of information on randomisation and assigned treatment (trandolapril or placebo). Adjustment for these covariates did not affect the results regarding the impact of age and sex. The analysis of trandolapril versus placebo showed a similar effect of age on the risk of SCD. The risk was non-SCD. Information about other medical treatment issued at discharge was not recorded and could not be analysed.

Adjustment for other risk factors likely to be associated with SCD or non-SCD risk and their relation did not change the results. Additional risk factors included in the model were wall motion index, congestive heart failure, New York Heart Association functional class, ventricular fibrillation, atrial fibrillation, smoking, body mass index, history of MI, hypertension, angina, or diabetes, and thrombolysis. The independent risk ratio associated with male sex was 1.25 (p < 0.05) for SCD and 1.12 (p = 0.2) for non-SCD. For a 10 year increase in age the risk ratios were 1.25 and 1.73, respectively (p < 0.0001 for both SCD and non-SCD). Furthermore, decreased wall motion index, congestive heart failure, New York Heart Association functional class, atrial fibrillation, and history of hypertension, angina, or diabetes were all independently associated with increased risk of both SCD and non-SCD.

### DISCUSSION

This is the first study of unselected and consecutive MI survivors to report the impact of age and sex on the risk of SCD and non-SCD. The principal findings of this study are that the risk of SCD increases with age, that the relation between risk of SCD and non-SCD decreases with age, and that male sex is associated with a slightly increased risk of SCD independent of age.

#### Impact of age

Our results are in perfect agreement with the results from the Framingham heart study, obtained from the general population during 38 years of follow up. Previous analyses of the subgroup with overt coronary heart disease did not show a clear effect of increasing age on the risk of SCD. The SCD rate was found to be similar regardless of age for men. The proportion of SCD relative to all cardiovascular death was found to decrease, in a rather unstable fashion, with increasing age but conclusions were limited because of the small number of women. Before the introduction of thrombolysis, post-MI studies in Sweden found a similar effect of age on the proportion of mortality being caused by SCD as we have reported here. The SCD rate was not given or tested in different age groups but seemed stable. No differences were found in the SCD risk between men and women. These studies included deaths during the initial hospitalisation, which may account for some of these differences.

In a study of 477 post-MI patients, Odemuyiwa and colleagues divided the patients into two age groups (below and above 60 years) and found a similar SCD mortality in the

<table>
<thead>
<tr>
<th>Age group (years)</th>
<th>&lt;56 (n=1120)</th>
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<th>66-75 (n=2001)</th>
<th>≥76 (n=1331)</th>
</tr>
</thead>
<tbody>
<tr>
<td>All cause death</td>
<td>105</td>
<td>273</td>
<td>623</td>
<td>658</td>
</tr>
<tr>
<td>Cardiovascular death</td>
<td>79</td>
<td>209</td>
<td>477</td>
<td>496</td>
</tr>
<tr>
<td>SCD</td>
<td>46</td>
<td>109</td>
<td>205</td>
<td>176</td>
</tr>
<tr>
<td>Non-SCD</td>
<td>33</td>
<td>100</td>
<td>272</td>
<td>320</td>
</tr>
<tr>
<td>Non-cardiovascular death</td>
<td>15</td>
<td>39</td>
<td>97</td>
<td>105</td>
</tr>
<tr>
<td>Unclassifiable death</td>
<td>11</td>
<td>25</td>
<td>49</td>
<td>57</td>
</tr>
</tbody>
</table>

SCD, sudden cardiovascular death.

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**Table 2** Frequency of different modes of death according to age group
two groups: 4.7% and 3.2%, respectively, and a significantly lower proportion of mortality from SCD in patients older than 60 years. In that study exclusion criteria were not having a Holter recording performed, age above 75 years, and any condition affecting heart rate variability other than ischaemic heart disease (that is, atrial fibrillation and diabetes mellitus). Owing to this selection the population studied had a relatively low mortality and the actual number of events was only 19 SCDs.

Many postinfarction studies have not classified death with respect to mode of death since their main purpose was to assess a change in all cause mortality between the different treatment groups. In the large β blocker trials the mode of death was examined in the population randomly assigned to treatment, but the data published have specifically addressed the issue of whether β blocker treatment reduces SCD. There have been no reports on the epidemiological distribution of SCD in these trials but, being part of randomised trials, patients are highly selected and not representative of the true postinfarction population.

Impact of sex
We found male sex to have an independent value in predicting SCD, while it had no influence on non-SCD risk. In the Framingham heart study male sex was found to increase the age adjusted SCD rate by a factor of four among patients with overt coronary heart disease. This may indicate that the absolute risk for women is so low that certain kinds of intervention would not be appropriate or cost effective. In contrast to this we found male sex to increase the risk of SCD by only 1.34 independently of age. This may result from differences between our population of MI survivors and the population with coronary heart disease in the Framingham heart study, defined as patients with angina pectoris and prior MI, including silent MI. An even more probable explanation may be the increase in female cardiovascular risk over time, which may be responsible for the observed difference, since the Framingham heart study reports results from 38 years of follow up. In the Framingham heart study the incidence of sudden death among women was similar to that among men 20 years younger. In our study the mortality caused by SCD among women in a given age group was always slightly higher than that of men in the younger age group. Thus, women lagged less than 10 years behind men with regard to the incidence of SCD.

Table 3

<table>
<thead>
<tr>
<th>Age (years)</th>
<th>SCD (%)</th>
<th>Non-SCD (%)</th>
<th>Ratio between hazards of SCD and non-SCD*</th>
</tr>
</thead>
<tbody>
<tr>
<td>≤56</td>
<td>4.8</td>
<td>3.4</td>
<td>1.44</td>
</tr>
<tr>
<td>56–65</td>
<td>7.3</td>
<td>6.9</td>
<td>1.09</td>
</tr>
<tr>
<td>66–75</td>
<td>10.5</td>
<td>13.7</td>
<td>0.76</td>
</tr>
<tr>
<td>&gt;76</td>
<td>14.2</td>
<td>25.3</td>
<td>0.55</td>
</tr>
</tbody>
</table>

*p<0.001 for no difference between age groups.

Table 4

<table>
<thead>
<tr>
<th>Age (years)</th>
<th>SCD (%)</th>
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<td>0.76</td>
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To study closely the influence of menopause it would be tempting to divide the youngest group but, as is evident from table 2, the numbers of events are already low, especially when stratified by sex. The effect of sex on SCD risk was not statistically different in different age groups, as would be expected if hormone concentrations conferred a major risk increase, but a severe limitation in that aspect is the limited number of events among young women. Furthermore, women who have established coronary heart disease are more likely to be deficient in
female sex hormones and this probably increases their cardiovascular risk.2,22 Thus, the women in our study are highly selected compared with women in the general population, probably also with respect to hormone concentrations. The only reported randomised mortality study of hormone replacement therapy (HERS (heart and oestrogen/progestin replacement study))21 enrolled women with established coronary heart disease and no benefit was seen. Primary prevention trials are ongoing.

Male sex was weakly but significantly associated with an increased risk of all cause mortality, independently of age. In contrast, a recent study found that female sex adds risk independently of age following MI.24 This probably relies on the observation that female sex is associated with increased in-hospital mortality during acute MI.25 The present study included exclusively survivors of MI.

The crude risk of sudden death

The cumulative incidence of SCD during the first year was somewhat higher than in CAMI (Canadian assessment of myocardial infarction),26 another epidemiological study among post-MI survivors in the thrombolytic era. CAMI reported an incidence of arrhythmic death during the initial year following MI survival of only 1.9%. The CAMI population was younger, since the initial 2477 patients were below 75 years of age. This is reflected in the all cause mortality rate (7.1%), which is less than half that of our population (14.9%) at one year. Furthermore, the CAMI study used the CAST (cardiac arrhythmia suppression trial) criteria27 for classification of arrhythmic death, which tends to give a lower estimate than SCD.28 Comparison with the MPIP (multicenter post-infarction program)29 is problematic since MPIP was conducted before thrombolysis in patients less than 70 years of age. Arrhythmic death was classified using the Hinkle and Thaler criteria.1

Pathophysiological considerations

Age related pathoanatomical changes include diffuse myocardial fibrosis and myocyte disarray, which may provide an anatomical substrate for arrhythmia.29 Advanced age is also associated with changes in autonomic nervous system activity with increased sympathetic and decreased vagal tone,29,30 which may facilitate arrhythmia in the presence of an anatomical substrate. However, myocardial fibrosis and increased sympathetic activity may also be causal factors in the development of heart failure. Such mechanisms may explain the increased incidence in our study of both SCD and non-SCD death with increasing age. On the other hand, increased vagal activation may accompany acute myocardial ischaemia in women more often than in men,31 which may be one reason for male sex being a risk factor for SCD.

Study limitations

SCD may not always be arrhythmic. However, in several studies—MADIT (multicenter automatic defibrillator implantation trial),4 AVID (antiarrhythmics versus implantable defibrillators), CIDS (Canadian implantable defibrillator study), and CASH (cardiac arrest study Hamburg)—implantable cardioverter defibrillators have been found to reduce in particular death classified as arrhythmic, giving an indication that classification is indeed meaningful. The result of classification of mode of death is highly dependent on the amount of data available,14 but the data needed to classify death as SCD are usually available from the death certificate, emergency room charts, or hospital files. However, we are aware of the inaccuracy of death certificates and have avoided any default classification. This is reflected in the 8.6% of cases where information was inadequate to classify the mode of death.

We have no data on the medical treatment assigned at discharge. The increased use of β blockers and statins is likely to reduce the absolute risk of both SCD and non-SCD. So far the effect of treatment has not been isolated to any particular age group or sex. This means that even if the absolute risk is reduced, the relative risk associated with age and sex remains the same despite changes in medical treatment. Similar considerations apply for the increased use of invasive treatments in post-MI patients.

Implications

The present study shows that, even if the proportion of cardiovascular mortality from SCD decreases with increasing age, the overall risk of SCD increases significantly. This implies that specific intervention against SCD should be tested or applied in all age groups of MI survivors. Furthermore, as the incidence of SCD is highest in the older age groups, the elderly will probably benefit the most from treatment as indexed by the CIDS trial.30 Effective treatment such as implantation of a cardioverter defibrillator should not be restricted to younger patients. Aggressive and effective treatment should be applied in the patients at highest risk, such as the elderly with a poor ejection fraction. The concept that these patients will die from other causes if SCD is prevented has never been found to be true.

Future interventional studies should include both men and women and investigate whether treatment effects are different between sexes. Since the crude SCD risk in men is only slightly higher than that in women, sex should not be considered when the indication for intervention is evaluated.

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REFERENCES

Near-fatal arrhythmia caused by hyperkalaemia

A 75 year old woman with end stage diabetic renal disease was found with bradycardia (15–20 beats per minute) and asystole. After cardiopulmonary resuscitation (CPR) and administration of atropine (2 mg) and ephedrine (100 μg), her heart rhythm converted to a stable rhythm, and she was transferred to the hospital. On admission, the ECG revealed a normal sinus rhythm with prolonged PQ interval, widening of QRS complexes, and peaked T waves (panel A). Forty minutes later, complete atrioventricular heart block developed with no ventricular escape rhythm (panel B) again necessitating CPR and administration of atropine and ephedrine. A subsequent ECG showed merging of QRS complexes with T waves (sine wave pattern) (panel C). At this time, the serum potassium concentration was 8.4 mmol/L, and blood glucose was raised (26.3 mmol/L). The patient was treated with calcium gluconate, insulin, and sodium bicarbonate as well as immediate haemodialysis. Potassium lowering treatment resulted in progressive narrowing of QRS complexes (panel D), serum potassium concentration 7.1 mmol/L (panel E), serum potassium concentration 6.3 mmol/L (panel F) within the next few hours. With a serum potassium of 4.4 mmol/L, her ECG showed no changes to previous ECGs (panel F).

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References

Near-fatal arrhythmia caused by hyperkalaemia

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